

# Estimation Methods with Ordered Covariate Subject to Measurement Error and Missingness in Semi-Ecological Design

by

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## Supplementary Material

### EM with Measurement Errors Only:

#### 1) Linear regression

First note that  $\theta = (\theta_1, \theta_2, \theta_3)$  where  $\theta_1 = (\beta_0, \beta_1, \sigma_\varepsilon^2)$ ,  $\theta_2 = \sigma_\eta^2$  and  $\theta_3 = (\mu, \sigma_b^2)$ . The E-Step of the  $t$ th iteration of the EM procedure gives

$$\begin{aligned} Q(\theta|\theta^{(t)}) &= E_{\theta^{(t)}}[l_c(\theta; \mathbf{Y}, \mathbf{W}, \mathbf{X})|\mathbf{y}, \mathbf{w}] \\ &= -\frac{1}{2} \left( \sum_{g=1}^G n_g \right) (\ln \sigma_\eta^2 + \ln \sigma_b^2 + \ln \sigma_\varepsilon^2) - \frac{1}{2\sigma_\varepsilon^2} \sum_{g=1}^G \sum_{i=1}^{n_g} E_{\theta^{(t)}}[(Y_{gi} - \beta_0 - \beta_1 X_{gi})^2 | y_{gi}, w_{gi}] \\ &\quad - \frac{1}{2\sigma_\eta^2} \sum_{g=1}^G \sum_{i=1}^{n_g} E_{\theta^{(t)}}[(W_{gi} - X_{gi})^2 | y_{gi}, w_{gi}] - \frac{1}{2\sigma_b^2} \sum_{g=1}^G \sum_{i=1}^{n_g} E_{\theta^{(t)}}[(X_{gi} - \mu_g)^2 | y_{gi}, w_{gi}] \end{aligned} \quad (1)$$

In M-Step, we need to maximize  $Q(\theta|\theta^{(t)})$  under the constraints  $\mu_1 \leq \mu_2 \leq \dots \leq \mu_G$ . For this, note first the conditional variable  $(X_{gi}|y_{gi}, w_{gi})$  follows  $N(m_x(y_{gi}, w_{gi}; \theta), v_x(\theta))$ , where

$$m_x(y_{gi}, w_{gi}; \theta) = \frac{\beta_1 \sigma_b^2 \sigma_\eta^2 (y_{gi} - \beta_0) + (\sigma_\varepsilon^2 \sigma_\eta^2) w_{gi} + \sigma_\varepsilon^2 \sigma_\eta^2 \mu_g}{\beta_1^2 \sigma_b^2 \sigma_\eta^2 + \sigma_\varepsilon^2 (\sigma_b^2 + \sigma_\eta^2)} \quad \text{and} \quad v_x(\theta) = \frac{\sigma_\varepsilon^2 \sigma_b^2 \sigma_\eta^2}{\beta_1^2 \sigma_b^2 \sigma_\eta^2 + \sigma_\varepsilon^2 (\sigma_b^2 + \sigma_\eta^2)}.$$

Let  $\bar{m}_g = \frac{1}{n_g} \sum_{i=1}^{n_g} m_x(y_{gi}, w_{gi}; \theta^{(t)})$ ,  $g = 1, \dots, G$ ,  $\bar{m} = \frac{1}{\sum_{g=1}^G n_g} \sum_{g=1}^G n_g \bar{m}_g$ , and  $\bar{y} = \frac{1}{\sum_{g=1}^G n_g} \sum_{g=1}^G \sum_{i=1}^{n_g} y_{gi}$ . Then, the solution to this maximization problem can be found and updated as follows:

$$\mu^{(t+1)} = \text{isotonic regression of } (\bar{m}_1, \bar{m}_2, \dots, \bar{m}_G)' \text{ with weight vector } (n_1, n_2, \dots, n_G)',$$

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$$\beta_1^{(t+1)} = \frac{\sum_{g=1}^G \sum_{i=1}^{n_g} [m_x(y_{gi}, w_{gi}; \theta^{(t)}) - \bar{m}] y_{gi}}{\sum_{g=1}^G \sum_{i=1}^{n_g} [m_x(y_{gi}, w_{gi}; \theta^{(t)}) - \bar{m}]^2},$$

$$\beta_0^{(t+1)} = \bar{y} - \beta_1^{(t+1)} \bar{m},$$

$$\sigma_b^2{}^{(t+1)} = \frac{1}{\sum_{g=1}^G n_g} \sum_{g=1}^G \sum_{i=1}^{n_g} [m_x(y_{gi}, w_{gi}; \theta^{(t)}) - \mu_g^{(t+1)}]^2 + v_x(\theta^{(t)}),$$

$$\sigma_\varepsilon^2{}^{(t+1)} = \frac{1}{\sum_{g=1}^G n_g} \sum_{g=1}^G \sum_{i=1}^{n_g} [y_{gi} - \beta_0^{(t+1)} - \beta_1^{(t+1)} m_x(y_{gi}, w_{gi}; \theta^{(t)})]^2 + v_x(\theta^{(t)}).$$

If we keep updating estimates by this EM algorithm, then  $\theta^{(t)}$  will converge the true MLE of  $\theta$ . Note that no Monte Carlo method is necessary for the simple linear case.

## 2) Logistic regression

Parameters in the logistic regression model are  $\theta_1 = (\beta_0, \beta_1)$ ,  $\theta_2 = \sigma_\eta^2$  and  $\theta_3 = (\mu, \sigma_b^2)$ . As in the simple linear case,  $\sigma_\eta^2$  is assumed to be known. The E step for this model gives

$$\begin{aligned} Q(\theta|\theta^{(t)}) &= E_{\theta^{(t)}}[l_c(\theta; \mathbf{Y}, \mathbf{W}, \mathbf{X})|\mathbf{y}, \mathbf{w}] \\ &= -\frac{1}{2} \left( \sum_{g=1}^G n_g \right) (\ln \sigma_\eta^2 + \ln \sigma_b^2) \\ &\quad + \sum_{g=1}^G \sum_{i=1}^{n_g} \{ y_{gi} E_{\theta^{(t)}}[\ln p(X_{gi}; \beta)|y_{gi}, w_{gi}] + (1 - y_{gi}) E_{\theta^{(t)}}[\ln(1 - p(X_{gi}; \beta))|y_{gi}, w_{gi}] \} \\ &\quad - \frac{1}{2\sigma_\eta^2} \sum_{g=1}^G \sum_{i=1}^{n_g} E_{\theta^{(t)}}[(W_{gi} - X_{gi})^2|y_{gi}, w_{gi}] - \frac{1}{2\sigma_b^2} \sum_{g=1}^G \sum_{i=1}^{n_g} E_{\theta^{(t)}}[(X_{gi} - \mu_g)^2|y_{gi}, w_{gi}] \end{aligned} \quad (2)$$

In fact, the third term of  $Q(\theta|\theta^{(t)})$  is constant because  $\sigma_\eta^2$  is known. Since the conditional density of  $X_{gi}$  given  $Y_{gi} = y_{gi}$  and  $W_{gi} = w_{gi}$  is

$$f(x_{gi}|y_{gi}, w_{gi}; \theta) = \frac{p(x_{gi}; \beta)^{y_{gi}} [1 - p(x_{gi}; \beta)]^{1-y_{gi}} h(x_{gi}; w_{gi}, \mu_i, \sigma_\eta^2, \sigma_b^2)}{\int p(x_{gi}; \beta)^{y_{gi}} [1 - p(x_{gi}; \beta)]^{1-y_{gi}} h(x_{gi}; w_{gi}, \mu_i, \sigma_\eta^2, \sigma_b^2) dx_{gi}}$$

where  $h(x_{gi}; w_{gi}, \mu_i, \sigma_b^2, \sigma_\eta^2)$  is the p.d.f of  $N(\frac{\sigma_b^2 w_{gi} + \sigma_\eta^2 \mu_i}{\sigma_b^2 + \sigma_\eta^2}, \frac{\sigma_b^2 \sigma_\eta^2}{\sigma_b^2 + \sigma_\eta^2})$ , conditional expectations in  $Q(\theta|\theta^{(t)})$  do not have closed form of expressions. Thus, a Monte-Carlo EM method is used as is generally the case in many similar situations. The outline of the M-Step in the  $(t+1)$ st iteration of the EM algorithm can be described as follows:

Step 1: Set  $\mu^{(t+1)}$  equal to the isotonic regression of  $(\bar{m}_1, \dots, \bar{m}_G)'$  with weight vector  $(n_1, \dots, n_G)'$ , where  $\bar{m}_g = \frac{1}{n_g} \sum_{i=1}^{n_g} E_{\theta^{(t)}}[X_{gi}|y_{gi}, w_{gi}]$ . Then, compute  $\sigma_b^2{}^{(t+1)} = \frac{1}{\sum_{g=1}^G n_g} \sum_{g=1}^G \sum_{i=1}^{n_g} E_{\theta^{(t)}}[(X_{gi} - \mu^{(t+1)})^2|y_{gi}, w_{gi}]$ , similarly.

Step 2: Keeping  $\theta^{(t)}$  in the conditional distribution, apply a usual Newton method to maximize  $Q(\theta|\theta^{(t)})$  with respect to  $\beta$  until a convergence criterion is satisfied. And set  $\beta^{(t+1)}$  equal to the solution.

It should be noted that the Newton method in Step 2 can be applied simply to the second term in  $Q(\theta|\theta^{(t)})$  because all other conditional expectations do not involve  $\beta$ .

## EM with Measurement Errors and Missing in Covariate:

### 1) Linear regression

In this case,  $Q_1(\theta|\theta^{(t)})$  is the same as (4) while  $Q_2(\theta|\theta^{(t)})$  is given by

$$\begin{aligned} Q_2(\theta|\theta^{(t)}) &= -\frac{1}{2} \left( \sum_{g=1}^G n_g^* \right) (\ln \sigma_\eta^2 + \ln \sigma_b^2 + \ln \sigma_\varepsilon^2) - \frac{1}{2\sigma_\varepsilon^2} \sum_{g=1}^G \sum_{i=1}^{n_g^*} E_{\theta^{(t)}} [(Y_{gi}^* - \beta_0 - \beta_1 X_{gi}^*)^2 | y_{gi}^*] \\ &\quad - \frac{1}{2\sigma_\eta^2} \sum_{g=1}^G \sum_{i=1}^{n_g^*} E_{\theta^{(t)}} [(W_{gi}^* - X_{gi}^*)^2 | y_{gi}^*] - \frac{1}{2\sigma_b^2} \sum_{g=1}^G \sum_{i=1}^{n_g^*} E_{\theta^{(t)}} [(X_{gi}^* - \mu_g)^2 | y_{gi}^*]. \end{aligned} \quad (3)$$

Recall first  $(X_{gi}|y_{gi}, w_{gi})$  follows  $N(m_x(y_{gi}, w_{gi}; \theta), v_x(\theta))$ . Also note that  $(X_{gi}^*, W_{gi}^* | y_{gi}^*)$  follows a bivariate normal distribution  $BVN(m_{x^*}(y_{gi}^*; \theta), m_{w^*}(y_{gi}^*; \theta), \rho_{x^*w^*}(\theta), v_{x^*}(\theta), v_{w^*}(\theta))$  where  $m_{x^*}(y_{gi}^*; \theta) = m_{w^*}(y_{gi}^*; \theta) = \frac{\sigma_\varepsilon^2 \mu_g + \beta_1 \sigma_b^2 (y_{gi}^* - \beta_0)}{\sigma_\varepsilon^2 + \beta_1^2 \sigma_b^2}$ ,  $\rho_{x^*w^*}(\theta) = \left[ \frac{\sigma_\varepsilon^2 \sigma_b^2}{\sigma_\varepsilon^2 \sigma_b^2 + \sigma_\eta^2 \sigma_{\eta 2}^2 + \beta_1^2 \sigma_b^2 \sigma_\eta^2} \right]^{1/2}$ ,  $v_{x^*}(\theta) = \frac{\sigma_\varepsilon^2 \sigma_b^2}{\sigma_\varepsilon^2 + \beta_1^2 \sigma_b^2}$ , and  $v_{w^*}(\theta) = \frac{\sigma_\varepsilon^2 \sigma_b^2 + \sigma_\eta^2 \sigma_{\eta 2}^2 + \beta_1^2 \sigma_b^2 \sigma_\eta^2}{\sigma_\varepsilon^2 + \beta_1^2 \sigma_b^2}$ .

Similarly to the case without missing, let  $\bar{m}_g = \frac{1}{n_g + n_g^*} \sum_{i=1}^{n_g} [m_x(y_{gi}, w_{gi}; \theta^{(t)}) + m_{x^*}(y_{gi}^*; \theta^{(t)})]$ ,  $g = 1, \dots, G$ ,  $\bar{m} = \frac{1}{\sum_{g=1}^G (n_g + n_g^*)} \sum_{g=1}^G (n_g + n_g^*) \bar{m}_g$ , and  $\bar{y} = \frac{1}{\sum_{g=1}^G (n_g + n_g^*)} \sum_{g=1}^G (\sum_{i=1}^{n_g} y_{gi} + \sum_{i=1}^{n_g^*} y_{gi}^*)$ . Then, considering (1) and (3), we can establish the EM algorithm that updates estimates as follows:

$$\begin{aligned} \mu^{(t+1)} &= \text{isotonic regression of } (\bar{m}_1, \bar{m}_2, \dots, \bar{m}_G)' \text{ with weight vector } (n_1 + n_1^*, n_2 + n_2^*, \dots, n_G + n_G^*)', \\ \beta_1^{(t+1)} &= \frac{\sum_{g=1}^G \{ \sum_{i=1}^{n_g} [m_x(y_{gi}, w_{gi}; \theta^{(t)}) - \bar{m}] y_{gi} + \sum_{i=1}^{n_g^*} [m_{x^*}(y_{gi}^*; \theta^{(t)}) - \bar{m}] y_{gi}^* \}}{\sum_{g=1}^G \{ \sum_{i=1}^{n_g} [m_x(y_{gi}, w_{gi}; \theta^{(t)}) - \bar{m}]^2 + \sum_{i=1}^{n_g^*} [m_{x^*}(y_{gi}^*; \theta^{(t)}) - \bar{m}]^2 \}}, \\ \beta_0^{(t+1)} &= \bar{y} - \beta_1^{(t+1)} \bar{m}, \\ \sigma_b^{2(t+1)} &= \frac{1}{\sum_{g=1}^G (n_g + n_g^*)} \sum_{g=1}^G \left\{ \sum_{i=1}^{n_g} [m_x(y_{gi}, w_{gi}; \theta^{(t)}) - \mu_g^{(t+1)}]^2 + [m_{x^*}(y_{gi}^*; \theta^{(t)}) - \mu_g^{(t+1)}]^2 \right. \\ &\quad \left. + n_g v_x(\theta^{(t)}) + n_g^* v_{x^*}(\theta^{(t)}) \right\}, \\ \sigma_\varepsilon^{2(t+1)} &= \frac{1}{\sum_{g=1}^G (n_g + n_g^*)} \sum_{g=1}^G \left\{ \sum_{i=1}^{n_g} [y_{gi} - \beta_0^{(t+1)} - \beta_1^{(t+1)} m_x(y_{gi}, w_{gi}; \theta^{(t)})]^2 \right. \\ &\quad \left. + \sum_{i=1}^{n_g^*} [y_{gi}^* - \beta_0^{(t+1)} - \beta_1^{(t+1)} m_{x^*}(y_{gi}^*; \theta^{(t)})]^2 + n_g v_x(\theta^{(t)}) + n_g^* v_{x^*}(\theta^{(t)}) \right\}. \end{aligned}$$

### 2) Logistic regression

Based on observations having missing values in covariate, the second term of  $Q(\theta|\theta^{(t)})$  for this model is

expressed as

$$\begin{aligned}
Q_2(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)}) &= -\frac{1}{2} \left( \sum_{g=1}^G n_g \right) (\ln \sigma_\eta^2 + \ln \sigma_b^2) \\
&+ \sum_{g=1}^G \sum_{i=1}^{n_g^*} \{ y_{gi}^* E_{\boldsymbol{\theta}^{(t)}} [\ln p(X_{gi}^*; \boldsymbol{\beta}) | y_{gi}^*] + (1 - y_{gi}^*) E_{\boldsymbol{\theta}^{(t)}} [\ln(1 - p(X_{gi}^*; \boldsymbol{\beta})) | y_{gi}^*] \} \\
&- \frac{1}{2\sigma_\eta^2} \sum_{g=1}^G \sum_{i=1}^{n_g^*} E_{\boldsymbol{\theta}^{(t)}} [(W_{gi}^* - X_{gi}^*)^2 | y_{gi}^*] - \frac{1}{2\sigma_b^2} \sum_{g=1}^G \sum_{i=1}^{n_g^*} E_{\boldsymbol{\theta}^{(t)}} [(X_{gi}^* - \mu_g)^2 | y_{gi}^*]. \tag{4}
\end{aligned}$$

In order to maximize  $Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)})$ , we need a Newton method as a part of each EM procedure. However, our investigations indicate that it does not take too long time to reach a convergence criterion. Considering (2) and (4), the M-Step can be summarized as follows:

Step 1: Set  $\boldsymbol{\mu}^{(t+1)}$  equal to the isotonic regression of  $(\bar{m}_1, \dots, \bar{m}_G)'$  with weight vector  $(n_1 + n_1^*, \dots, n_G + n_G^*)'$ , where  $\bar{m}_g = \frac{1}{n_g + n_g^*} \{ \sum_{i=1}^{n_g} E_{\boldsymbol{\theta}^{(t)}} [X_{gi} | y_{gi}, w_{gi}] + \sum_{i=1}^{n_g^*} E_{\boldsymbol{\theta}^{(t)}} [X_{gi}^* | y_{gi}^*] \}$ . Then, compute  $\sigma_b^{2(t+1)} = \frac{1}{\sum_{g=1}^G (n_g + n_g^*)} \sum_{g=1}^G \{ \sum_{i=1}^{n_g} E_{\boldsymbol{\theta}^{(t)}} [(X_{gi} - \mu^{(t+1)})^2 | y_{gi}, w_{gi}] + \sum_{i=1}^{n_g^*} E_{\boldsymbol{\theta}^{(t)}} [(X_{gi}^* - \mu^{(t+1)})^2 | y_{gi}^*] \}$ , similarly.

Step 2: Keeping  $\boldsymbol{\theta}^{(t)}$  in the conditional distributions, plug  $\boldsymbol{\mu}^{(t+1)}$  and  $\sigma_b^{2(t+1)}$  into  $Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)})$  and apply a usual Newton method to maximize  $Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)})$  with respect to  $\boldsymbol{\beta}$ . Set  $\boldsymbol{\beta}^{(t+1)}$  equal to the solution.

As mentioned earlier, the conditional expectations here do not have closed form of expressions, and thus we rely on a Monte Carlo method to evaluate them.